

The background of the slide features a large, faint watermark of the University of Vienna seal. The seal is circular and contains a central figure, likely a seated scholar or saint, surrounded by Latin text and various heraldic symbols like eagles and shields.

# Probability Theory

## 14. Poisson convergence

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## Example: Binomial $\Rightarrow$ Poisson

- ▶ Let  $p_1, p_2, \dots \in [0, 1]$  such that  $np_n \rightarrow \lambda \geq 0$ .

Then,  $B(n, p_n) \xrightarrow{n \rightarrow \infty} \text{Poi}(\lambda)$ .

Indeed:

$$\begin{aligned}\psi_{B(n, p_n)}(t) &= \left(1 - p_n(1 - e^{it})\right)^n \\ &= \left(1 - \frac{n \cdot p_n}{n}(1 - e^{it})\right)^n \\ &\xrightarrow{n \rightarrow \infty} \exp(-\lambda(1 - e^{it})) = \psi_{\text{Poi}(\lambda)}(t).\end{aligned}$$

# The generating function

- ▶ Let  $X$  be rv with values in  $\mathbb{Z}_+$ . Then,

$$z \mapsto \varphi_X(z) := P[z^X] = \sum_{k=0}^{\infty} z^k P(X = k)$$

is called generation function (of the distribution) of  $X$ . With

$$z = e^{-t},$$

$$\mathcal{L}_X(t) = P[e^{-tX}] = P[z^X] = \varphi_X(z).$$

- ▶ Generating function is distribution-determining;

Weak convergence  $\iff$  Convergence of the gener. fcts.;

- ▶ Note that

$$\varphi'_X(1) = \sum_{k=0}^{\infty} k z^{k-1} P(X = k) \Big|_{z=1} = \sum_{k=0}^{\infty} k P(X = k) = P[X].$$

## Asymptotic negligibility

- ▶ Definition 10.4: A family  $(X_{nj})_{n=1,2,\dots,n,j=1,\dots,m_n}$  with  $m_1, m_2, \dots \in \mathbb{N}$  is *asymptotically negligible* if  $X_{n1}, \dots, X_{n,m_n}$  is independent,  $n = 1, 2, \dots$  and for all  $\varepsilon > 0$

$$\sup_{j=1,\dots,m_n} P(|X_{nj}| > \varepsilon) \xrightarrow{n \rightarrow \infty} 0.$$

If  $X_{ij} \geq 0$  for all  $i, j$ , then  $m_n = \infty$  is also allowed.

- ▶ For  $\mathbb{Z}$ -valued rvs this is equivalent to

$$\left. \begin{array}{l} \inf_{j=1,\dots,m_n} P(|X_{nj}| = 0) \\ \inf_{j=1,\dots,m_n} E[|X_{nj}| \wedge 1] \\ \inf_{j=1,\dots,m_n} \varphi_{X_{nj}}(0) \end{array} \right\} \xrightarrow{n \rightarrow \infty} 1.$$

## A lemma

- Lemma 10.6:  $(\lambda_{nj})_{n=1,2,\dots,j=1,\dots,m_n}$  non-negative,  $\lambda \geq 0$ . Then

$$\prod_{j=1}^{m_n} (1 - \lambda_{nj}) \xrightarrow{n \rightarrow \infty} e^{-\lambda} \quad \iff \quad \sum_{j=1}^{m_n} \lambda_{nj} \xrightarrow{n \rightarrow \infty} \lambda.$$

Proof:  $\log(1 - x) = -x + o(x)$ . LHS equivalent to

$$\begin{aligned} -\lambda &= \lim_{n \rightarrow \infty} \sum_{j=1}^{m_n} \log(1 - \lambda_{nj}) = - \lim_{n \rightarrow \infty} \sum_{j=1}^{m_n} \lambda_{nj} \left(1 - \frac{\varepsilon(\lambda_{nj})}{\lambda_{nj}}\right) \\ &= - \lim_{n \rightarrow \infty} \sum_{j=1}^{m_n} \lambda_{nj}. \end{aligned}$$

## Poisson convergence

- Theorem 10.5:  $(X_{nj})_{n=1,2,\dots,n,j=1,\dots,m_n}$  asymptotically negligible,  $\mathbb{Z}_+$ -valued,  $X \sim \text{Poi}(\lambda)$ . Then

$$\sum_{j=1}^{m_n} X_{nj} \xrightarrow{n \rightarrow \infty} X \iff \left( \begin{array}{l} \sum_{j=1}^{m_n} \mathbb{P}(X_{nj} > 1) \xrightarrow{n \rightarrow \infty} 0, \\ \sum_{j=1}^{m_n} \mathbb{P}(X_{nj} = 1) \xrightarrow{n \rightarrow \infty} \lambda. \end{array} \right)$$

$\Leftarrow$ :  $\varphi_{n,j} := \varphi_{X_{n,j}}$ , to show  $\prod_{j=1}^{m_n} \varphi_{nj}(z) \xrightarrow{n \rightarrow \infty} e^{-\lambda(1-z)}$  or

$$A_n(z) := \sum_{j=1}^{m_n} (1 - \varphi_{nj}(z)) \xrightarrow{n \rightarrow \infty} \lambda(1-z),$$

We write

$$\begin{aligned} A_n(z) &= \sum_{j=1}^{m_n} 1 - \mathbb{P}(X_{nj} = 0) - z\mathbb{P}(X_{nj} = 1) + o(1) \\ &= \sum_{j=1}^{m_n} (1-z)\mathbb{P}(X_{nj} = 1) \xrightarrow{n \rightarrow \infty} \lambda(1-z). \end{aligned}$$

## Poisson convergence of geometrically distributed rv

- ▶  $Y_{nj} + 1 \sim \text{geo}(p_n)$ ,  $j = 1, \dots, n$ ,  $n = 1, 2, \dots$ . We set

$$Y_n := \sum_{j=1}^n Y_{nj},$$

Number of failures before the  $n$ th success.

If  $Y \sim \text{Poi}(\lambda)$  and  $(1 - p_n) \cdot n \xrightarrow{n \rightarrow \infty} \lambda$ , then  $Y_n \xrightarrow{n \rightarrow \infty} Y$ .

Indeed:

$$\sum_{j=1}^n \mathbb{P}(Y_{nj} = 1) = n(1 - p_n)p_n \xrightarrow{n \rightarrow \infty} \lambda,$$

$$\sum_{j=1}^n \mathbb{P}(Y_{nj} > 1) = n(1 - p_n)^2 \xrightarrow{n \rightarrow \infty} 0$$